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**POST-DISCHARGE OUTCOMES ARE SIMILAR FOR WEEKEND VERSUS WEEKDAY DISCHARGES FOR GENERAL INTERNAL MEDICINE PATIENTS ADMITTED TO TEACHING HOSPITALS**

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**Conflict of Interest Disclosures:** None **Abstract**

**Background:** Hospitals reduce staffing levels and services on weekends. This raises the question of whether weekend discharges may be inadequately prepared and thus at higher risk for adverse events post-discharge.

**Objective:** To comparedeath or non-elective readmission rates 30-days after weekend vs. weekday discharge.

**Design:** Retrospective cohort.

**Setting:** All teaching hospitals in Alberta, Canada.

**Patients:** General Internal Medicine (GIM) discharges (only one per patient).

**Measurements:** Analyses were adjusted for demographics, comorbidity, and length of stay based on a previously validated index.

**Results**: Of 7991 patients (mean age 62.1, 51.9% male, mean Charlson 2.56, 57.5% LACE >=10) discharged from 7 teaching hospitals, 1146 (14.3%) were discharged on a weekend. Although they had substantially shorter lengths of stay (5.64 days [95%CI 5.35-5.93] vs. 7.86 days [95%CI 7.71-8.00], adjusted p-value < 0.0001) and were less likely to be discharged with homecare support (10.9% vs. 19.3%) or to long-term care facilities (3.1% vs. 7.8%), patients discharged on weekends exhibited similar rates of death or readmission at 30-days compared to those discharged on weekdays (10.6% vs. 13.2%, aOR 0.94, 95%CI 0.77-1.16), even among the 4591 patients deemed to be high risk for post-discharge events based on LACE score ≥10: 16.8% versus 16.5% for weekday discharges, aOR 1.09 (95%CI 0.85-1.41).

**Conclusions:** GIM patients discharged from teaching hospitals on weekends have shorter lengths of stay and exhibit similar post-discharge outcomes as patients discharged on weekdays.

Hospitals typically reduce staffing levels and the availability of diagnostic, laboratory, and treatment services on weekends, and patients ***admitted*** on weekends exhibit poorer in-hospital outcomes for several medical conditions.[1-9] Whether or not patients ***discharged*** on weekends have worse clinical outcomes has been less well studied.[10-12] Discharge rates on Saturday and Sunday are lower than for the other five days of the week,[12] but bed shortages and hospital overcrowding have increased the demand for maximizing 24-7 week-round discharge efficiency. Given that the number of patients discharged on weekends is likely to continue to increase, it is important to assess the risk of weekend discharge on outcomes monitored as performance indicators by organizations such as the Centers for Medicare and Medicaid Services, the American Medical Association Physicians Consortium for Performance Improvement, the National Quality Forum, and the Joint Commission.

Thus, we designed this study to evaluate baseline characteristics, length of stay, and post-discharge outcomes for General Internal Medicine patients in teaching hospitals discharged on weekends compared to weekdays. Our objective was to determine whether post-discharge outcomes differed for patients discharged on weekends vs. weekdays.

**Methods:**

***Study Setting:***

The Canadian province of Alberta has a single vertically integrated health care system that is government-funded and provides universal access to hospitals, Emergency Departments (ED), and outpatient physician services for all 4.1 million Albertans as well as all prescription medications for the poor, socially disadvantaged, disabled, or those 65-years and older. This study received approval from the University of Alberta Health Research Ethics Board with waiver of informed consent.

***Data Sources:***

This study used de-identified linked data from 3 Alberta Health administrative databases that capture vital status and all hospital or ED visits and have previously been shown to have high accuracy for medical diagnoses.[13] The Alberta Health Care Insurance Plan Registry tracks date of death or emigration from the province. The Discharge Abstract Database includes the most responsible diagnosis identified by the hospital attending physician, up to 25 other diagnoses coded by nosologists in each hospital, the admission and discharge dates, the admission category (elective or urgent/emergent) for all acute care hospitalizations. Of note, unlike US studies, the hospital databases are able to distinguish in-hospital (e.g., adverse events) versus pre-morbid diagnoses (e.g., pre-existing comorbidities). The Ambulatory Care Database captures all patient visits to Emergency Departments with coding for up to 10 conditions per encounter.

***Study Cohort:***

We identified all adults with an acute care hospitalization on the General Internal Medicine (GIM) services at all 7 Alberta teaching hospitals (i.e., defined as those with Royal College of Physicians and Surgeons of Canada-approved residency training programs in internal medicine, the equivalent of the AAMC certification in the US) between October 1, 2009 and September 30, 2010 and between April 1, 2011 and December 1, 2011 (these 20 months covered most of the pre/post intervals for a recently reported quality improvement initiative at one of the teaching hospitals which had no significant impact on post-discharge outcomes).[14] Patients from out of province or transferred from/to another inpatient service (e.g., the intensive care unit, a different service in the same hospital [such as surgery], another acute care hospital, or rehabilitation hospital) or with lengths of stay greater than 30 days were excluded. We only included the first hospitalization for any one patient in our study timeframe and thus excluded repeat discharges of the same patient.

***Explanatory Variable of Interest:***

 The independent variable of interest was calendar day of discharge, stratified according to weekday (Monday thru Friday) vs weekend (Saturday and Sunday). Only 1.4% of weekday discharges occurred on a statutory holiday and for the purposes of this study these discharges were also considered weekend discharges. At the 7 teaching hospitals in Alberta, nursing staffing ratios do not differ between weekend and weekday, but availability of all other members of the health care team does. Physician census decreases from 4-5 per ward to 1-2, and ward-based social workers, occupational therapists, physiotherapists, and pharmacist educators are generally not available on weekends.

***Outcomes:***

 Our primary outcome of interest was the composite outcome of death or all-cause non-elective readmission within 30 days of discharge (ie. not including in-hospital events prior to discharge or elective readmissions after discharge for planned procedures such as chemotherapy); hereafter we refer to this as “death or readmission.” This is a patient-relevant outcome that is highlighted in the Affordable Care Act and for which there are several validated risk adjustment models.[15] We chose a composite outcome to deal with the issue of competing risks (if weekend discharges were more likely to die then we could observe a spurious association between weekend discharge and reduced readmissions if we focused on only that outcome).

***Other Measures:***

Comorbidities for each patient were identified using ICD-9 and ICD-10 codes from the Discharge Abstract Database for the index hospitalization and any hospitalizations in the 12 months prior to their index admission, a method previously validated in Alberta databases.[13] We also recorded health resource use during their index hospitalization and calculated each patient’s LACE score at the time of discharge, an index for predicting unplanned readmission or early death post discharge previously validated in Canadian administrative databases.[15] The LACE index includes length of hospital stay (“L”), acuity of admission (“A”, based on the admission category variable described earlier), comorbidity burden quantified using the Charlson comorbidity index (“C”), and emergency department visits in the six months prior to admission (“E”); patients with discharge LACE scores greater than 10 (total possible score is 19) are defined as being at “high risk” of death/readmission within 30 days.[16] As detailed below, in order to deal with potential concerns that LOS may be a mediator in the causal pathway, we ran two sensitivity analyses: one in which we excluded LOS from the analyses and one in which we included expected LOS rather than the actual LOS: expected LOS are data driven estimates based on the most current 2 years of patient LOS information available in the Canadian Institute for Health Information discharge abstract database (www.cihi.ca) for all acute care hospitals in Canada and was generated for each patient independently of our study taking into account case mix group, age, and inpatient resource intensity weights.

***Statistical Analysis***

Baseline patient characteristics between weekend and weekday discharges were compared with t-tests for continuous variables and X2 tests for binary or categorical variables. Logistic regression was used for comparison of death or readmission for weekend vs weekday discharges. Multivariable models were adjusted for age, sex, hospital and LACE scores (as a continuous variable) at time of discharge; in sensitivity analyses we adjusted for (i) LACE score without including LOS and (ii) LACE score using expected LOS rather than actual LOS. In further sensitivity analyses we (i) restricted the analysis to only those patients deemed to be at high risk for events due to LACE scores of 10 or greater and (ii) included ED visits as part of the composite endpoint (i.e., death, unplanned readmission, or unplanned ED visit within 30 days of discharge). Day of admission (weekend vs weekday) was also considered for the multivariable models but was not found to be significant and thus was omitted from final models. We do not have any physician identifying variables in our dataset and thus could not investigate the potential correlation among patients discharged by the same physician. We did explore the hospital intra-class correlation coefficient and as it was very small (0.001) we did not utilize models to account for the hierarchical nature of the data but did include hospital as a fixed effect in the logistic models (results were virtually identical whether we did or did not include hospital in the models). Adjusted odds ratios (ORs) are displayed with 95% confidence intervals (CI) and p-values. Average length of stay was calculated for weekend and weekday discharges with 95% confidence intervals. P-values for adjusted length of stay were calculated using multivariable linear regression adjusting for age, sex, day of admission and Charlson score. All statistical analyses were done using SAS for Windows version 9.4 (Cary, NC).

**RESULTS**

***Patient Characteristics***

Of the 7991 patients discharged during our study interval, 1146 (14.3%) were discharged on weekends or holiday days (Table 1) – in contrast, 2180 of our cohort were admitted on a weekend (27.3%). The mean age of our study population was 62.1 years of age, 51.9% were men, mean Charlson score was 2.56, and 4591 (57.5%) had LACE scores of at least 10 at discharge.

***Weekday vs Weekend Discharge***

Although patients admitted on weekdays and weekends were very similar (data available upon request), patients discharged on weekends (compared to those discharged on weekdays) were younger, more likely to be discharged home without additional supports, and had fewer comorbidities (Table 1, Figure 1). Patients discharged on weekends had shorter lengths of stay than those discharged on weekdays (5.6 days vs 7.9 days, p<0.0001). In adjusted linear regression analyses, this 2.3 day difference remained statistically significant (adjusted p-value <0.0001).

Patients discharged on a weekend exhibited lower unadjusted 30-day rates of death or readmission than those discharged on a weekday (10.6% vs 13.2%), but these differences disappeared after multivariable adjustment that accounted for differences in risk profile: aOR 0.94, 95% CI 0.77 to 1.16 (Table 2). Results were similar in sensitivity analyses adjusting for LACE scores without LOS included (aOR 0.88, 95% CI 0.71 to 1.08) or adjusting for LACE scores using expected LOS rather than actual LOS (aOR 0.90, 95% CI 0.73 to 1.10). Restricting the analysis to only those patients deemed to be at high risk for events due to LACE scores of 10 or greater confirmed that weekend and weekday discharges had similar outcomes in the first 30 days after discharge (aOR 1.09, 95% CI 0.85-1.41, Table 2). Similar patterns were seen when we included ED visits as part of the composite endpoint (i.e., death, unplanned readmission, or unplanned ED visit within 30 days of discharge, Table 2).

**DISCUSSION**

Our data suggests that patients discharged from the GIM teaching wards we studied on weekends were appropriately triaged as they did not exhibit a higher risk of adverse events post-discharge. Although patients discharged on weekends tended to be younger and had less comorbidities than those discharged during the week, we adjusted for baseline covariates in analyses and we did not find an association between weekend discharge and increased post-discharge events even amongst the subset of patients deemed to be at high risk for post-discharge adverse events (based on high LACE scores). To our knowledge, although we previously examined this issue in patients with a most responsible diagnosis of heart failure,[10] examining weekend vs weekday discharges in the full gamut of general medical patients admitted to teaching hospitals has not previously been examined.

In our previous study[10] of over 24,000 heart failure patients discharged over 10 years (up to June 2009, so no overlap with any patients in this study) we also found that patients discharged on the weekends were younger, had less comorbidities, and shorter lengths of stay. Although post-discharge death/readmission rates were higher for weekend discharged patients in our earlier study (21.1% vs. 19.5%, aHR 1.15, 95%CI 1.06-1.25), it is worth noting that this was almost entirely driven by data from ***non-teaching hospitals and cardiology wards***. Thus, it is important to reiterate that the findings in our current study are for GIM wards in teaching hospitals and may not be generalizable to less structured non-teaching settings.

Although we did not study physician decision-making, our results suggest that physicians are incorporating discharge day into their discharge decision-making. They may be selecting younger patients with less comorbidities for weekend discharges, or they may be delaying the discharges of older patients with more comorbidities for weekday discharges. Either is not surprising given the realities of weekend inpatient care: reduced staffing and frequent cross-coverage (of physicians, nurses and other allied health care workers), limited support services (such as laboratory services or diagnostic imaging), and decreased availability of community services (including home care and social support services).[17] For example, in one large US heart failure registry, patients discharged on a weekend received less complete discharge instructions than those discharged on weekdays.[11] Given that early follow-up post-discharge is associated with better outcomes,[18,19] future studies should also explore whether patterns of patient follow-up differ after weekend vs. weekday discharges.

Although we were able to capture all interactions with the health care system in a single payer system with universal access, there are some limitations to our study. First, we used administrative data, which precludes fully adjusting for severity of diagnoses or functional status although we used proxies such as admission from/discharge to long term care facility.[20,21] Second, we did not have access to process of care measures such as diagnostic testing or prescribing data and thus cannot determine whether quality of care or patient adherence differed by day of the week they were discharged on, although this seems unlikely. Third, although post-discharge follow-up may be associated with better outcomes,[18,19] we were unable to adjust for patterns of outpatient follow-up in this study. Fourth, we acknowledge that death or readmission soon after discharge does not necessarily mean that the quality of care during the preceding hospitalization was suboptimal or that these deaths or readmissions were even potentially preventable. Many factors influence post-discharge mortality and/or readmission and quality of inpatient care is only one.[22-25] Fifth, although some may express concern that LOS may be a mediator in the causal pathway between discharge decision and post-discharge events and that adjusting for LOS in analyses could thus spuriously obscure a true association, it is worth pointing out that our two sensitivity analyses to explore this (the one in which we excluded LOS from the analyses and the one in which we included expected LOS rather than the actual LOS) revealed nearly identical point estimates and 95% CI as our main analysis. Finally, as our study is observational we cannot definitively conclude causality. Nor can we exclude an 18% excess risk for patients discharged on weekends (or a 22% lower risk either), given our 95% CI for post-discharge adverse outcomes.

**Conclusion:**

 We found that the proportion of patients discharged on weekends is lower than the proportion admitted on weekends. We also found that lower risk / less severely ill patients appear to be preferentially discharged on weekends, and as a result post-discharge outcomes are similar between weekend and weekday discharges despite shorter LOS and less availability of outpatient resources for weekend discharged patients. The reasons why more complicated patients are not discharged on weekends deserves further study as safely increasing weekend discharge rates would improve efficiency and safety (by reducing unnecessary exposure to in-hospital adverse events such as falls, unnecessary urinary catheterizations, and healthcare-acquired infections). While admission to hospital has become a 24-7 business, we believe that hospital discharge processes should strive for the same level of efficiency.

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**Table 1.** Characteristics of General Internal Medicine patients discharged from 7 teaching hospitals

|  |  |  |  |
| --- | --- | --- | --- |
|  | **Weekend Discharge** | **Weekday Discharge** | **P-value** |
| No. of patients | 1146 | 6845 |  |
| Age, mean (SD) | 57.97 (19.70) | 62.77 (19.37) | <.0001 |
| Male | 601 (52.4) | 3548 (51.8) | 0.70 |
| Top 5 most responsible diagnoses |  |  |  |
| COPD | 74 (6.5) | 507 (7.4) |  |
| Pneumonia | 64 (5.6) | 326 (4.8) |  |
| Heart failure | 31 (2.7) | 375 (5.5) |  |
| Urinary Tract Infection | 39 (3.4) | 254 (3.7) |  |
| Venous thromboembolism | 31 (2.7) | 259 (3.8) |  |
| Charlson score, mean (SD) | 2.17 (3.29) | 2.63 (3.30) | <.0001 |
| Comorbidities (based on index hospitalization and prior 12 months) |  |
| Hypertension | 485 (42.3) | 3265 (47.7) | 0.00 |
| Diabetes mellitus | 326 (28.4) | 2106 (30.8) | 0.11 |
| Fluid imbalance | 332 (29.0) | 1969 (28.8) | 0.89 |
| COPD | 255 (22.3) | 1790 (26.2) | 0.01 |
| Psychiatric disorder | 179 (15.6) | 1459 (21.3) | <.0001 |
| Pneumonia | 242 (21.1) | 1427 (20.8) | 0.84 |
| Anemia | 167 (14.6) | 1233 (18.0) | 0.00 |
| Trauma | 169 (14.7) | 1209 (17.7) | 0.02 |
| Atrial fibrillation | 141 (12.3) | 1069 (15.6) | 0.00 |
| Heart Failure | 101 (8.8) | 946 (13.8) | <.0001 |
| Drug abuse | 188 (16.4) | 966 (14.1) | 0.04 |
| Cancer | 124 (10.8) | 867 (12.7) | 0.08 |
| Renal disease | 93 (8.1) | 689 (10.1) | 0.04 |
| Dementia | 49 (4.3) | 564 (8.2) | <.0001 |
| Mild liver disease | 99 (8.6) | 587 (8.6) | 0.94 |
| Cerebrovascular disease | 59 (5.1) | 492 (7.2) | 0.01 |
| Gastrointestinal bleed | 84 (7.3) | 496 (7.2) | 0.92 |
| Asthma | 83 (7.2) | 426 (6.2) | 0.19 |
| Stroke | 42 (3.7) | 332 (4.9) | 0.08 |
| Prior myocardial infarction | 47 (4.1) | 329 (4.8) | 0.30 |
| Arthritis | 42 (3.7) | 309 (4.5) | 0.19 |
| Peripheral vascular disease | 42 (3.7) | 259 (3.8) | 0.84 |
| Severe liver disease | 44 (3.8) | 261 (3.8) | 0.97 |
| Valve disease | 24 (2.1) | 188 (2.7) | 0.20 |
| Paralysis | 31 (2.7) | 201 (2.9) | 0.67 |
| Skin ulcer | 17 (1.5) | 137 (2.0) | 0.24 |
| Shock | 19 (1.7) | 99 (1.4) | 0.58 |
| HIV | 15 (1.3) | 109 (1.6) | 0.47 |
| Protein calorie malnutrition | 0 (0.0) | 9 (0.1) | 0.21 |
| Features of index hospitalization |  |  |  |
| Resource intensity weight, mean (SD) | 1.10 (0.82) | 1.38 (1.24) | <.0001 |
| LACE score, mean (SD) | 9.45 (2.85) | 10.51 (3.03) | <.0001 |
| Expected LOS, mean (SD) | 6.20 (4.08) | 7.12 (4.89) | <.0001 |
| Acute LOS, mean (SD) | 5.64 (4.99) | 7.86 (6.13) | <.0001 |
| Weekend Admission | 244 (21.3) | 1936 (28.3) | <.0001 |
| Discharge Disposition |  |  | <.0001 |
| Transferred to another inpatient hospital  | 14 (1.2) | 189 (2.8) |  |
| Transferred to Long Term Care facility | 36 (3.1) | 532 (7.8) |  |
| Transferred to other (ex. hospice) | 5 (0.4) | 24 (0.4) |  |
| Discharged to home setting with support services | 125 (10.9) | 1318 (19.3) |  |
| Discharged home | 926 (80.8) | 4646 (67.9) |  |
| Left against medical advice | 40 (3.5) | 136 (2.0) |   |

Numbers are n (%) unless specified otherwise

**Table 2.** Post-discharge Outcomes after a General Internal Medicine hospitalization in a teaching hospital †

|  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- |
| **Death / readmission within 30 days** | **Weekend discharge n/N (%)** | **Weekday discharge n/N (%)** | **Unadjusted p-value** | **aOR**† **(95% CI)** | **Adjusted p-value** |
| All 7 teaching hospitals, all pts | 121/1146 (10.6) | 901/6845 (13.2) | 0.01 | 0.94 (0.77, 1.16) | 0.58 |
| All 7 teaching hospitals, but only pts with LACE < 10 | 37/647 (5.7) | 225/2753 (8.2) | 0.04 | 0.72 (0.50, 1.03) | 0.07 |
| All 7 teaching hospitals, but only pts with LACE >= 10 | 84/499 (16.8) | 676/4092 (16.5) | 0.86 | 1.09 (0.85, 1.41) | 0.49 |
| **Death / readmission / ED visit within 30 days** |  |  |  |  |  |
| All 7 teaching hospitals, all pts | 218/1146 (19.0) | 1445/6845 (21.1) | 0.11 | 0.98 (0.83, 1.15) | 0.79 |
| All 7 teaching hospitals, but only pts with LACE < 10 | 90/647 (13.9) | 460/2753 (16.7) | 0.08 | 0.83 (0.64, 1.06) | 0.13 |
| All 7 teaching hospitals, but only pts with LACE >= 10 | 128/499 (25.7) | 985/4092 (24.1) | 0.44 | 1.12 (0.90, 1.39) | 0.31 |
| **Death within 30 days** |  |  |  |  |  |
| All 7 teaching hospitals, all pts | 24/1146 (2.1) | 215/6845 (3.1) | 0.05 | 0.97 (0.63, 1.51) | 0.89 |
| All 7 teaching hospitals, but only pts with LACE < 10 | 4/647 (0.6) | 23/2753 (0.8) | 0.58 | 0.89 (0.30, 2.62) | 0.83 |
| All 7 teaching hospitals, but only pts with LACE >= 10 | 20/499 (4.0) | 192/4092 (4.7) | 0.49 | 0.99 (0.61, 1.61) | 0.98 |
| **Readmission within 30 days** |  |  |  |  |  |
| All 7 teaching hospitals, all pts | 105/1146 (9.2) | 751/6845 (11.0) | 0.07 | 0.94 (0.76, 1.17) | 0.59 |
| All 7 teaching hospitals, but only pts with LACE < 10 | 33/647 (5.1) | 211/2753 (7.7) | 0.02 | 0.68 (0.46, 0.99) | 0.04 |
| All 7 teaching hospitals, but only pts with LACE >= 10 | 72/499 (14.4) | 540/4092 (13.2) | 0.44 | 1.14 (0.87, 1.49) | 0.34 |
| **ED visit within 30 days** |  |  |  |  |  |
| All 7 teaching hospitals, all pts | 182/1146 (15.9) | 1118/6845 (16.3) | 0.70 | 1.00 (0.84, 1.19) | 0.99 |
| All 7 teaching hospitals, but only pts with LACE < 10 | 83/647 (12.8) | 412/2753 (15.0) | 0.17 | 0.84 (0.65, 1.09) | 0.20 |
| All 7 teaching hospitals, but only pts with LACE >= 10 | 99/499 (19.8) | 706/4092 (17.3) | 0.15 | 1.17 (0.92, 1.48) | 0.20 |

 † Multivariable models adjust for age, sex, hospital, and LACE score at time of discharge from index hospitalization. Weekday discharge is reference group for odds ratios.

